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\* Part of the slides were made when I was an  
Advanced Algorithmic Engineer at Trip.com

# Conditional Quantile Regression

With Applications to User-Preferred Price Prediction

December 23, 2021

# Table of Contents

[Introduction](#)

[Methodology: Quantile Regression](#)

[Offline Evaluations](#)

[Discussion and Future Works](#)

# Introduction to Our Hotel Ranking Task

A group of  
candidate hotels (in  
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\*Logistic Regression, XGBoost, Deep Neural Networks,...

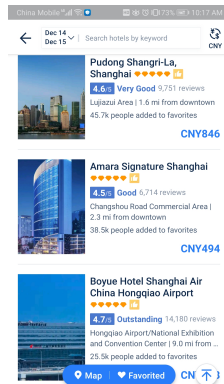
# Introduction to Our Hotel Ranking Task

A group of candidate hotels (in a searched city).



Ranking Algorithms\*

A well-sorted list of hotels.



\*Logistic Regression, XGBoost, Deep Neural Networks,...

## Objective of the Hotel Ranking Task

Return a list of hotels with user-preferred ones placed on the top.

⇒ Optimizing the *conversion rate* (on hotels with high commissions).



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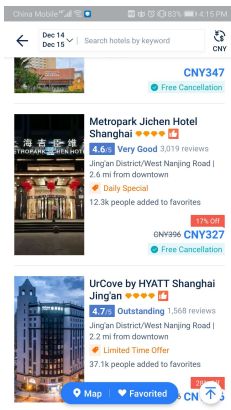
Features/Predictors:  $\mathbf{X}_i = \left[ \underbrace{V_1^{(i)}, \dots, V_q^{(i)}}_{\text{Hotel Features}}, \underbrace{U_1^{(i)}, \dots, U_p^{(i)}}_{\text{User Features}} \right]$  for  $i = 1, \dots, n$ .

Responses:  $Y_i \in \{0 : \text{Not Booked}, 1 : \text{Booked}\}$  for  $i = 1, \dots, n$ .

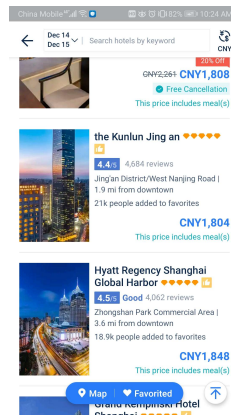
## How to Identify User-Preferred Hotels?

- The prices of hotels clicked/booked by a user quantify his/her affordability.
- The price preferences of users on our platform are diverse.

# Variety of User Price Preferences



(a) Users that prefer low-priced hotels

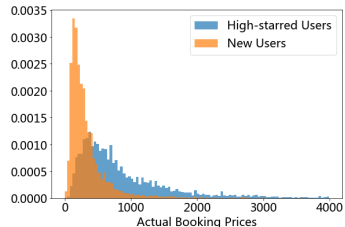
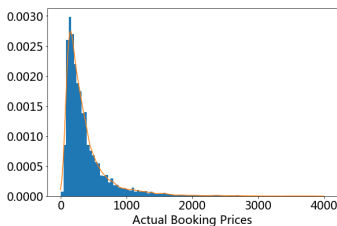


(b) Users that prefer high-priced hotels



## Multimodal Nature of User Price Preferences

The price preferences varies between different groups of users on our platform.



**Figure 2:** Overall and group-specific distributions of actual booking prices on December 6, 2021.

## Main Objective: User-Preferred Hotel Price Prediction

Correctly predicting the preferred hotel prices or *price intervals* is of great significance to our hotel ranking task!

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Mathematically, given a user  $\mathbf{X}_i = \mathbf{x}_i = [u_1^{(i)}, \dots, u_p^{(i)}]$ , we intend to predict his/her preferred price interval

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Here, the features  $u_j^{(i)}, j = 1, \dots, p$  range from

- user behaviors (such as historical clicked/booked hotels, user IDs, etc.)
- location information (such as city IDs, average GMV in that city, etc.)

## Drawback of the Current Online Model (Baseline)

**Current online model:** It is a weighted sum of historical booked prices, real-time clicked prices, and the specific quantile price in the searched city.

$$\text{Predicted Price} = \frac{\sum_i \omega_{\text{time}} \cdot \omega_{\text{type}} \cdot \omega_{\text{abnormal}} \cdot \omega_{\text{city}} \cdot \text{Price}_i}{\sum_i \omega_{\text{time}} \cdot \omega_{\text{type}} \cdot \omega_{\text{abnormal}} \cdot \omega_{\text{city}}}.$$

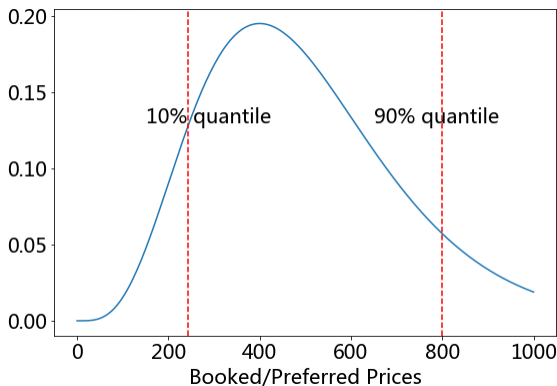
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- The choices weights  $\omega_{\text{time}}, \omega_{\text{type}}, \omega_{\text{abnormal}}, \omega_{\text{city}}$  are heuristic and outdated.
- The preferred price interval is symmetrically extended from the above point estimate.
- The accuracy of the current predicted prices (or price intervals) is also limited.
- ...

## Our Proposed Method: Conditional Quantile Regression



**Figure 3:** (Smoothed) conditional distribution of historical booked/preferred prices for a user with feature  $\mathbf{X}_i = \mathbf{x}_i$ . The synthetic density function (blue curve) is given by  $f(u|\mathbf{x}_i) = \frac{1}{\Gamma(5) \cdot 100^5} \cdot u^4 \exp\left(-\frac{u}{100}\right)$ .



## Our Proposed Method: Conditional Quantile Regression

Given the conditional cumulative distribution function  $F(y|\mathbf{X} = \mathbf{x})$  of booked prices, we pursue an interval

$$\left[ Q_{\tau}(\mathbf{x}), Q_{1-\tau}(\mathbf{x}) \right],$$

where  $Q_{\tau}(\mathbf{x}) = \inf \{y : F(y|\mathbf{X} = \mathbf{x}) \geq \tau\}$  and  $\tau \in (0, 1/2]$ .<sup>†</sup>

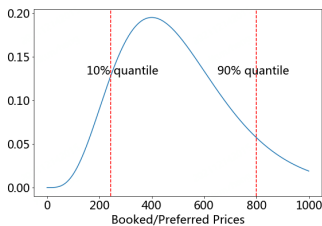


Figure 4:  $\tau$  and  $(1 - \tau)$  quantile of  $F(y|\mathbf{X} = \mathbf{x})$  with  $\tau = 0.1$ .

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<sup>†</sup>Koenker, R., & Bassett Jr, G. (1978). Regression quantiles. *Econometrica: Journal of the Econometric Society*, 33-50.

# How to Fit the Conditional Quantile?

## How to Fit the Conditional Quantile?

The conditional quantile  $Q_\tau(\mathbf{x})$  is the solution to the following optimization problem:

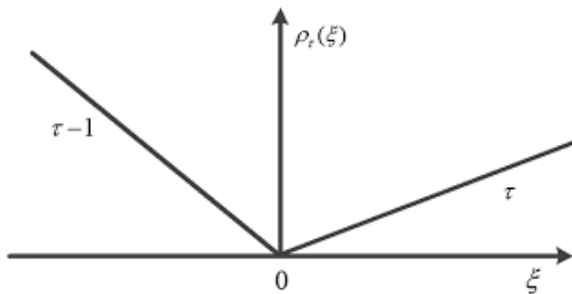
$$Q_\tau(\mathbf{x}) = \arg \min_q \mathbb{E} [\rho_\tau(Y - q) | \mathbf{X} = \mathbf{x}], \quad (1)$$

where

$$\rho_\tau(\xi) = \xi [\tau - \mathbb{1}_{\{\xi < 0\}}] = \begin{cases} \tau\xi, & \xi \geq 0, \\ -(1 - \tau)\xi, & \xi < 0 \end{cases} \quad (2)$$

is the so-called “pinball” loss (Koenker and Bassett, 1978; Firpo et al., 2009; Steinwart and Christmann, 2011).

## “Pinball Loss”



### Remark:

- When  $\tau = 0.5$ , the aforementioned optimization problem (1) recovers the absolute deviation problem.
- The loss is robust to outliers (Hampel, 1971; John, 2015).

## Correctness of the (Conditional) Quantile Regression

### *Proposition*

*Given the conditional distribution function  $F(y|\mathbf{X} = \mathbf{x})$ ,*

$$Q_\tau(\mathbf{x}) = \inf \{y : F(y|\mathbf{X} = \mathbf{x}) \geq \tau\}$$

*is the solution to (1).*

*More generally, given any càdlàg function  $F(y)$ ,*

$$q_\tau = \inf \{y : F(y) \geq \tau\}$$

*is the solution to the unconditional quantile regression problem  $\arg \min_q \mathbb{E} [\rho_\tau(Y - q)]$ .*

## Quantile Regression in Practice

Theoretically,  $Q_\tau(\mathbf{x}) = \arg \min_q \mathbb{E} [\rho_\tau(Y - q) | \mathbf{X} = \mathbf{x}]$ .

## Quantile Regression in Practice

Theoretically,  $Q_\tau(\mathbf{x}) = \arg \min_q \mathbb{E} [\rho_\tau(Y - q) | \mathbf{X} = \mathbf{x}]$ .

Practically, given the training set with clicked/booked hotel entries

$$\{(\mathbf{X}_i, Y_i)\} = \left\{ \left( \left[ U_1^{(i)}, \dots, U_p^{(i)} \right], Y_i \right) \right\},$$

we solve the following empirical risk minimization (ERM) problem:

$$\hat{Q}_\tau = \arg \min_{Q \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^n \rho_\tau(Y_i - Q(\mathbf{X}_i)),$$

where  $\mathcal{F}$  is the function class spanned by our (neural network) models.

## Fitting the Empirical Quantiles $\hat{Q}_\tau$ and $\hat{Q}_{1-\tau}$

**Input:**  $\{(\mathbf{X}_i, Y_i)\} = \left\{ \left( \left[ U_1^{(i)}, \dots, U_p^{(i)} \right], Y_i \right) \right\}$ , where the continuous features are standardized and discrete ones are converted to embedding vectors.



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**Architecture:** One shared hidden layer  $512 \times 200$  with additional separate  $200 \times 100 \times 1$  full-connected Relu layers.

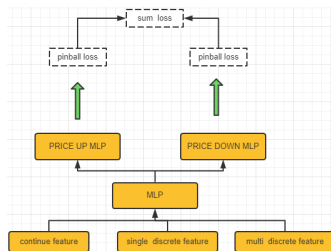


Figure 5: Double-tower architecture (image credit: Xianzhang Xiang)

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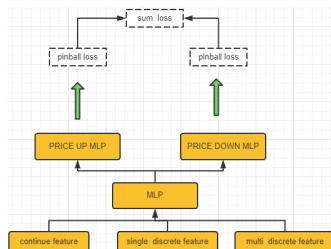


Figure 5: Double-tower architecture (image credit: Xianzhang Xiang)

**Objective:**  $\left\{ \hat{Q}_\tau, \hat{Q}_{1-\tau} \right\} = \arg \min_{\{f, g\} \subset \mathcal{F}} \frac{1}{n} \sum_{i=1}^n \left[ \rho_\tau(Y_i - f(\mathbf{X}_i)) + \rho_{1-\tau}(Y_i - g(\mathbf{X}_i)) \right]$

with  $\tau = 0.1$ .

## Why do we use Relu Neural Network? (Minimax Theory)

Assume that

- the true quantile function  $Q_\tau$  belongs to the Hölder class  $\mathcal{H}$  or Besov space  $\mathcal{B}$ .
- the number of layers  $L$  satisfies  $\log_2(n) \lesssim L \lesssim n^{\frac{p}{2s+p}}$ .
- the maximum norm of network coefficients  $\|\beta\|_{\max} \lesssim n^{\frac{p}{2s+p}} \log n$ .

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Then,

$$\|\widehat{Q}_\tau - Q_\tau\|_{\ell_2}^2 \leq C \cdot (\log n)^2 n^{-\frac{2s}{2s+p}},$$

where  $s$  is the smoothness parameter,  $p$  is the dimension of the feature space, and  $n$  is the sample size (Schmidt-Hieber, 2020; Padilla et al., 2020).

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Based on the nonparametric theory (Wasserman, 2006; Tsybakov, 2008), this rate of convergence is indeed *minimax* up to a log factor!

$\widehat{f}^*$  is minimax

$$\iff \sup_{f \in \mathcal{H}} \mathbb{E} \left[ \left( \widehat{f}^*(\mathbf{x}_0) - f(\mathbf{x}_0) \right)^2 \right] = \inf_{\widehat{f}_n} \sup_{f \in \mathcal{H}} \mathbb{E} \left[ \left( \widehat{f}_n(\mathbf{x}_0) - f(\mathbf{x}_0) \right)^2 \right],$$

where the infimum is taken among all the estimators.

## Summary of Our Proposed Model

- **Goal:** Preferred Price Interval  $\left[Q_\tau(\mathbf{x}), Q_{1-\tau}(\mathbf{x})\right]$  with  $Q_\tau(\mathbf{x}) = \inf \{y : F(y|\mathbf{X} = \mathbf{x}) \geq \tau\}$  and  $\tau \in (0, 1/2]$ .
- **Theoretical Solution:** Conditional Quantile Regression,

$$Q_\tau(\mathbf{x}) = \arg \min_q \mathbb{E} [\rho_\tau(Y - q) | \mathbf{X} = \mathbf{x}] .$$

- **Practical Model:** Empirical Risk Minimization with Relu Networks,

$$\hat{Q}_\tau = \arg \min_{Q \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^n \rho_\tau(Y_i - Q(\mathbf{X}_i)) .$$

- **Minimax Guarantee:**  $\|\hat{Q}_\tau - Q_\tau\|_{\ell_2}^2 \rightarrow 0$  as  $n \rightarrow \infty$ .

## Other Potential Choices of Quantile Regression Models

- **Quantile Regression Forests** (Meinshausen, 2006): The random forests method has the uniform consistency in estimating the cumulative distribution function (CDF) of  $Y|X = x$ .

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- **Quadratic Programming and Reproducing Kernel Hilbert Space (RKHS) Methods** (Takeuchi et al., 2006), **Nadaraya-Watson Nonparametric Regression Estimator** (Huang and Nguyen, 2018), etc.

## Evaluation Metrics

- **Coverage Accuracy:**

$$ACC(\mathcal{Y}, \hat{\mathcal{I}}) = \frac{1}{n} \sum_{i=1}^n \mathbb{1}_{\{Y_i \in [\hat{I}(x_i)]\}},$$

where  $\mathcal{Y} = \{Y_i\}_{i=1}^n$  is a collection of booked hotel prices and  $\hat{I}(x_i) = [\hat{Q}_\tau(x_i), \hat{Q}_{1-\tau}(x_i)]$  is the predicted preferred price interval for the user with feature  $x_i$ .

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- Average Interval Length:**

$$\text{Average Length}(\hat{\mathcal{I}}) = \frac{1}{n} \sum_{i=1}^n \left| \hat{Q}_\tau(x_i) - \hat{Q}_{1-\tau}(x_i) \right|.$$

## Neural Network Quantile Regression on the “My Location” Scenario

	Cov. Acc. ( <i>fh_prices</i> )	Cov. Acc. (or- der prices)	Average Inter- val Length
<b>Baseline Model</b>	0.7521	0.8019	283.8624
<b>Our NN QR</b>	<b>0.8718</b>	<b>0.8593</b>	<b>233.5811</b>

**Table 1:** Comparison between our neural network quantile regression model and the current online model (baseline) on the “My Location” scenario.

## Neural Network Quantile Regression on the “Main Ranking” Scenario

	Cov. Acc. (fh_prices)	Cov. Acc. (order prices)	Average Interval Length
<b>Baseline Interval I</b>	0.5590	0.5804	213.1197
<b>Baseline Interval II</b>	0.8593	0.8534	597.1224
<b>Our NN QR (Before calibration)</b>	0.7883	0.7351	317.5999
<b>Our NN QR (After calibration)</b>	<b>0.9268</b>	<b>0.8954</b>	482.3805

**Table 2:** Comparison between our neural network quantile regression model and the current online model (baseline) on the “Main Ranking” scenario.

- Notes: The calibration means that we extend our predicted interval as:

$$\left[ \hat{Q}_{\tau}(\mathbf{x}_i) - \alpha \cdot \left| \hat{Q}_{\tau}(\mathbf{x}_i) - \hat{Q}_{1-\tau}(\mathbf{x}_i) \right|, \hat{Q}_{1-\tau}(\mathbf{x}_i) + \alpha \cdot \left| \hat{Q}_{\tau}(\mathbf{x}_i) - \hat{Q}_{1-\tau}(\mathbf{x}_i) \right| \right],$$

where  $\alpha = 0.3 \sim 0.5$ .

## Discussion: Non-Crossing Property of Quantile Regression

Recall that our current optimization framework is

$$\{\hat{Q}_\tau, \hat{Q}_{1-\tau}\} = \arg \min_{\{Q_\tau, Q_{1-\tau}\} \subset \mathcal{F}} \frac{1}{n} \sum_{i=1}^n \left[ \rho_\tau(Y_i - Q_\tau(\mathbf{X}_i)) + \rho_{1-\tau}(Y_i - Q_{1-\tau}(\mathbf{X}_i)) \right].$$

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However, a constraint is required for the monotonicity of quantiles, *i.e.*, for any  $\tau \in (0, 1/2]$ , we should solve the constrained optimization problem:

$$\begin{aligned} \{\hat{Q}_\tau, \hat{Q}_{1-\tau}\} = \arg \min_{\{Q_\tau, Q_{1-\tau}\} \subset \mathcal{F}} \frac{1}{n} \sum_{i=1}^n \left[ \rho_\tau(Y_i - Q_\tau(\mathbf{X}_i)) + \rho_{1-\tau}(Y_i - Q_{1-\tau}(\mathbf{X}_i)) \right] \\ \text{subject to } Q_\tau(\mathbf{X}_i) \leq Q_{1-\tau}(\mathbf{X}_i) \text{ for all } i = 1, \dots, n. \end{aligned}$$

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**Challenges:** Solving the constrained optimization problem is difficult due to the nature of stochastic gradient descent (Padilla et al., 2020).



## Discussion: Solution to the Non-Crossing Constrained Quantile Regression

### Feasible Approaches:

- Penalized Method: With a large  $\lambda > 0$ , we optimize the following problem:

$$\begin{aligned} \left\{ \hat{Q}_\tau, \hat{Q}_{1-\tau} \right\} = \arg \min_{\{Q_\tau, Q_{1-\tau}\} \subset \mathcal{F}} \sum_{i=1}^n & \left[ \rho_\tau(Y_i - Q_\tau(\mathbf{X}_i)) + \rho_{1-\tau}(Y_i - Q_{1-\tau}(\mathbf{X}_i)) \right] \\ & + \lambda \cdot \sum_{i=1}^n \mathbb{1}_{\{\rho_\tau(Y_i - Q_\tau(\mathbf{X}_i)) > \rho_{1-\tau}(Y_i - Q_{1-\tau}(\mathbf{X}_i))\}}. \end{aligned}$$

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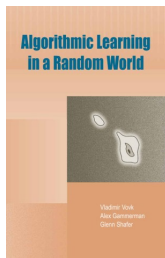
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- Redefined Objective (Padilla et al., 2020):

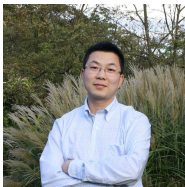
$$\left\{ \hat{h}_1, \hat{h}_2 \right\} = \arg \min_{\{h_1, h_2\} \subset \mathcal{F}} \sum_{i=1}^n \rho_\tau(Y_i - h_1(\mathbf{X}_i)) + \sum_{i=1}^n \rho_{1-\tau} \left\{ Y_i - h_1(\mathbf{X}_i) - \log \left[ 1 + e^{h_2(\mathbf{X}_i)} \right] \right\}$$

and set  $\hat{Q}_\tau(\mathbf{x}) = \hat{h}_1(\mathbf{x})$  and  $\hat{Q}_{1-\tau}(\mathbf{x}) = \hat{h}_1(\mathbf{x}) + \log \left[ 1 + e^{\hat{h}_2(\mathbf{x})} \right]$ .

# Motivation of Our Proposed Method: Conformal Inference



**Figure 6:** Algorithmic Learning in a Random World (Vovk et al., 2005).



**(a)** Jing Lei



**(b)** Larry Wasserman



**(c)** Emmanuel Candès

## Motivation of Our Proposed Method: Conformal Inference

What is conformal prediction/inference (Vovk et al., 1999, 2005; Lei et al., 2018)?

- Given a training set  $\{(\mathbf{X}_i, Y_i)\} \subset \mathbb{R}^p \times \mathbb{R}$  and the unknown value  $Y_{n+1}$  at a test point  $\mathbf{X}_{n+1}$ , it aims to construct a *marginal distribution-free prediction interval*  $\mathcal{C}(\mathbf{X}_{n+1}) \subset \mathbb{R}$  such that

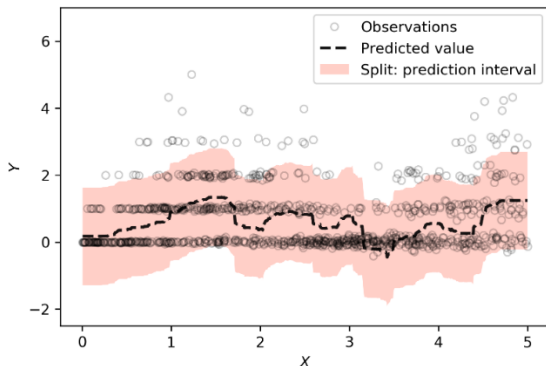
$$\mathbb{P}(Y_{n+1} \in \mathcal{C}(\mathbf{X}_{n+1})) \geq 1 - \alpha$$

for some nominal miscoverage level  $\alpha \in (0, 1)$ .

- Notes: The  $(1 - \alpha)$ -confidence interval is defined as:

$$\mathbb{P}(\mathbb{E}[Y|\mathbf{X}] \in \mathcal{C}(\mathbf{X})) \geq 1 - \alpha.$$

# Classical (Split) Conformal Prediction: A Preview



**Figure 8:** Classical (Split) Conformal Prediction (Average coverage: 91.4%; Average interval length: 2.91.)

## Classical (Split) Conformal Prediction: Detailed Procedures

- 1 Split the training set  $\mathcal{D} = \{(\mathbf{X}_i, Y_i)\} \subset \mathbb{R}^p \times \mathbb{R}$  into  $\mathcal{D} = \mathcal{D}_T \cup \mathcal{D}_C$ :
  - A proper training set  $\mathcal{D}_T = \{(\mathbf{X}_i, Y_i) : i \in \mathcal{I}_1\}$ ,
  - A calibration set  $\mathcal{D}_C = \{(\mathbf{X}_i, Y_i) : i \in \mathcal{I}_2\}$ .

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- 2 Fit  $\hat{\mu}(x) \leftarrow \mathcal{A}(\{(\mathbf{X}_i, Y_i) : i \in \mathcal{I}_1\})$  via any regression algorithm  $\mathcal{A}$  on  $\mathcal{D}_T$ .

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- 3 Compute the absolute residuals on  $\mathcal{D}_C$  as:

$$R_i = |Y_i - \hat{\mu}(\mathbf{X}_i)| \quad \text{with} \quad i \in \mathcal{I}_2.$$



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- 3 Compute the absolute residuals on  $\mathcal{D}_C$  as:

$$R_i = |Y_i - \hat{\mu}(\mathbf{X}_i)| \quad \text{with} \quad i \in \mathcal{I}_2.$$

- 4 Compute the  $(1 - \alpha)$  empirical quantile of the absolute residuals,  
 $Q_{1-\alpha}(R, \mathcal{I}_2) := (1 - \alpha) \left(1 + \frac{1}{|\mathcal{I}_2|}\right)$ -th empirical quantile of  $\{R_i : i \in \mathcal{I}_2\}$ .

## Classical (Split) Conformal Prediction: Detailed Procedures

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- 2 Fit  $\hat{\mu}(\mathbf{x}) \leftarrow \mathcal{A}(\{(\mathbf{X}_i, Y_i) : i \in \mathcal{I}_1\})$  via any regression algorithm  $\mathcal{A}$  on  $\mathcal{D}_T$ .
- 3 Compute the absolute residuals on  $\mathcal{D}_C$  as:

$$R_i = |Y_i - \hat{\mu}(\mathbf{X}_i)| \quad \text{with} \quad i \in \mathcal{I}_2.$$

- 4 Compute the  $(1 - \alpha)$  empirical quantile of the absolute residuals,
 
$$Q_{1-\alpha}(R, \mathcal{I}_2) := (1-\alpha) \left(1 + \frac{1}{|\mathcal{I}_2|}\right)\text{-th empirical quantile of } \{R_i : i \in \mathcal{I}_2\}.$$
- 5 The prediction interval at a new point  $\mathbf{X}_{n+1}$  is given by
 
$$\mathcal{C}(\mathbf{X}_{n+1}) = [\hat{\mu}(\mathbf{X}_{n+1}) - Q_{1-\alpha}(R, \mathcal{I}_2), \hat{\mu}(\mathbf{X}_{n+1}) + Q_{1-\alpha}(R, \mathcal{I}_2)].$$

## Conformalized Quantile Regression (Romano et al., 2019)

- 1 Split the training set  $\mathcal{D} = \{(\mathbf{X}_i, Y_i)\} \subset \mathbb{R}^p \times \mathbb{R}$  into  $\mathcal{D} = \mathcal{D}_T \cup \mathcal{D}_C$ :
  - A proper training set  $\mathcal{D}_T = \{(\mathbf{X}_i, Y_i) : i \in \mathcal{I}_1\}$ ,
  - A calibration set  $\mathcal{D}_C = \{(\mathbf{X}_i, Y_i) : i \in \mathcal{I}_2\}$ .
- 2 Fit  $\{\hat{Q}_{\alpha_{\text{low}}}, \hat{Q}_{\alpha_{\text{high}}}\} \leftarrow \mathcal{A}_q(\{(\mathbf{X}_i, Y_i) : i \in \mathcal{I}_1\})$  via any **quantile regression** algorithm  $\mathcal{A}_q$  on  $\mathcal{D}_T$ .

- 3 Compute the **conformity scores** of  $\hat{\mathcal{C}}(x) = [\hat{Q}_{\alpha_{\text{low}}}(x), \hat{Q}_{\alpha_{\text{high}}}(x)]$  on  $\mathcal{D}_C$  as:

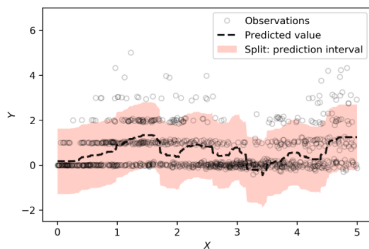
$$E_i := \max \left\{ \hat{Q}_{\alpha_{\text{low}}}(\mathbf{X}_i) - Y_i, Y_i - \hat{Q}_{\alpha_{\text{high}}}(\mathbf{X}_i) \right\} \quad \text{with} \quad i \in \mathcal{I}_2.$$

- 4 Compute the  $(1 - \alpha)$  empirical quantile of the conformity scores,
 
$$Q_{1-\alpha}(E, \mathcal{I}_2) := (1 - \alpha) \left( 1 + \frac{1}{|\mathcal{I}_2|} \right)\text{-th empirical quantile of } \{E_i : i \in \mathcal{I}_2\}.$$

- 5 The prediction interval at a new point  $\mathbf{X}_{n+1}$  is given by

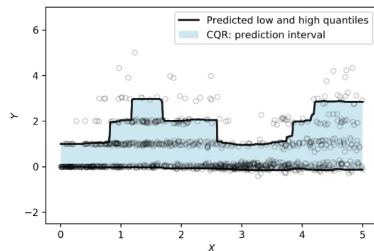
$$\mathcal{C}(\mathbf{X}_{n+1}) = \left[ \hat{Q}_{\alpha_{\text{low}}}(\mathbf{X}_{n+1}) - Q_{1-\alpha}(R, \mathcal{I}_2), \hat{Q}_{\alpha_{\text{high}}}(\mathbf{X}_{n+1}) + Q_{1-\alpha}(R, \mathcal{I}_2) \right].$$

## Comparisons Between Split Conformal Prediction and Conformalized Quantile Regression



(a) Classical (Split) Conformal Prediction

(Average coverage: 91.4%; Average interval length: 2.91).



(b) Conformalized Quantile Regression

(Average coverage: 91.06%; Average interval length: 1.99).

## Conclusion and Future Works

What we have done:

- We proposed a user-preferred price prediction model via (conditional) quantile regression with a Relu neural network.
- The model is well-performed based on offline evaluations.

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- We proposed a user-preferred price prediction model via (conditional) quantile regression with a Relu neural network.
- The model is well-performed based on offline evaluations.

Ongoing works:

- Handle the non-crossing properties/constraints of our model.
- Extend the user-preferred price prediction model to other scenarios and develop an unified modeling framework.

# Thank You

Comments or Questions?

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### Proof of Proposition 1.

Let  $g(u) = \mathbb{E} [\rho_\tau(Y - u)]$ . Some simple algebra show that

$$\begin{aligned} g(u) &= \int_{-\infty}^{\infty} \rho_\tau(y - u) dF(y) \\ &= \int_u^{\infty} \tau(y - u) dF(y) - \int_{-\infty}^u (1 - \tau)(y - u) dF(y). \end{aligned}$$

Applying the Leibniz integral rule shows that

$$g'(u) = 0 \iff -\tau \int_u^{\infty} dF(y) + (1 - \tau) \int_{-\infty}^u dF(y) = F(u) - \tau = 0.$$

Therefore,  $u = q_\tau$  is the smallest point satisfying  $F(u) - \tau = 0$  and will be unique when  $F$  is strictly monotonic on  $q_\tau$ .  $\square$